Empirical evaluation of fundamental principles of evidence-based medicine: a meta-epidemiological study

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Abstract

Rationale, aims and objectives 39 Evidence-based medicine (EBM) holds that estimates of effects of health interventions based on 40 high-certainty evidence (CoE) are expected to change less frequently than the effects generated in low CoE studies. However, this foundational principle of EBM has never been empirically 42 tested. 43 Methods 44 We reviewed all systematic reviews and meta-analyses in Cochrane Database of Systematic 45 Reviews from January 2016 through May 2021 (n=3,323). We identified 414(207x2) and 384 46 (192x2) pairs of original and updated Cochrane reviews that assessed CoE and pooled 47 treatment effect estimates. We appraised CoE using the Grading of Recommendations 48 Assessment, Development and Evaluation (GRADE) method. We assessed the difference in 49 effect sizes between the original versus updated reviews as a function of change in CoE, which 50 we report as a ratio of odds ratio (ROR). We compared ROR generated in the studies that 51 changed CoE from very low/low (VL/L) to moderate/high (M/H) vs. M/H ?VL/L. We also 52 assessed the heterogeneity and inconsistency (using the tau and I² statistic), the change in 53 precision of effect estimates (by calculating the ratio of standard errors) (seR), and the absolute 54 deviation in estimates of treatment effects (aROR). 55 Results 56 57 We found that CoE originally appraised as VL/L had 2.1 (95%CI: 1.19 to 4.12; p=0.0091) times 58 higher odds to be changed in the future studies than M/H CoE. However, the effect size was not 59 different when CoE changed from VL/L ?M/H vs. M/H ?VL/L [ROR=1.02 (95%CI: 0.74 to 1.39) 60 vs. 1.02 (95%CI: 0.44 to 2.37); p=1 for the between subgroup differences]. aROR was 61 similar 62 between the subgroups [median (IQR):1.12 (1.07 to 1.57) vs 1.21 (1.12 to 2.43)]. We observed 63 large inconsistency (12=99%) and imprecision in treatment effects (seR=1.09). 64 Conclusions 65 We provide the first empirical support for a foundational principle of EBM showing that low66 quality evidence changes more often than high CoE. However, the effect size was not different 67 between studies with low vs high CoE. The finding that the effect size did not differ between low 68 and high CoE indicate urgent need to refine current EBM critical appraisal methods.

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Evidence-based medicine (EBM) holds that estimates of effects of health interventions based on high-certainty evidence (CoE) are expected to change less frequently than the effects generated in low CoE studies. However, this foundational principle of EBM has never been empirically tested.

Methods
We reviewed all systematic reviews and meta-analyses in Cochrane Database of Systematic Reviews from January 2016 through May 2021 (n=3,323). We identified 414(207x2) and 384 (192x2) pairs of original and updated Cochrane reviews that assessed CoE and pooled treatment effect estimates. We appraised CoE
using the Grading of Recommendations Assessment, Development and Evaluation (GRADE) method. We assessed the difference in effect sizes between the original versus updated reviews as a function of change in CoE, which we report as a ratio of odds ratio (ROR). We compared ROR generated in the studies that changed CoE from very low/low (VL/L) to moderate/high (M/H) vs. MH/HVL/L. We also assessed the heterogeneity and inconsistency (using the tau and $I^2$ statistic), the change in precision of effect estimates (by calculating the ratio of standard errors) (seR), and the absolute deviation in estimates of treatment effects (aROR).

Results

We found that CoE originally appraised as VL/L had 2.1 (95%CI: 1.19 to 4.12; p=0.0091) times higher odds to be changed in the future studies than M/H CoE. However, the effect size was not different when CoE changed from VL/LM/H vs. M/HVL/L [ROR=1.02 (95%CI: 0.74 to 1.39) vs. 1.02 (95%CI: 0.44 to 2.37); p=1 for the between subgroup differences]. aROR was similar between the subgroups [median (IQR):1.12 (1.07 to 1.57) vs 1.21 (1.12 to 2.43)]. We observed large inconsistency ($I^2=99\%$) and imprecision in treatment effects (seR=1.09).

Conclusions

We provide the first empirical support for a foundational principle of EBM showing that low-quality evidence changes more often than high CoE. However, the effect size was not different between studies with low vs high CoE. The finding that the effect size did not differ between low and high CoE indicate urgent need to refine current EBM critical appraisal methods.

Key words: evidence-based medicine- critical appraisal-bias- random error- randomized trials - observational studies- systematic review- meta-epidemiology

Introduction

A foundational epistemological principle underpinning evidence-based medicine (EBM) is based on the assumption that the estimates of the effects of health interventions are closer to the “truth” if they are based on higher than on lower quality (certainty) of evidence (CoE).\(^1\) If the estimated treatment effects are close to the “true” effects, this would also imply that they would less likely to change as evidence accumulates after new studies are completed. Conversely, because its relation to the “truth” is less certain, this also implies that the estimated effects when evidence is of low quality would more likely change in future research. Research to date indicates that guideline panels are willing to issue stronger recommendations when they deem evidence to be of high quality, thus indirectly affirming this central EBM assumption.\(^2-5\)

However, whether this indirect assessment of quality of evidence based on guidelines panels’ decision-making is accurate is not known. It is possible that current methods of critical appraisal of CoE do not discriminate well between “true” accurate from inaccurate estimates of treatment effects. That is, the effects of health interventions based on low quality of evidence may turn out to reflect “true effects” by testing in subsequent studies. On the other hand, what was originally deemed as high quality evidence may be undermined by future studies more often than initially expected. Thus, it is not known if low quality evidence is more often revised than high quality evidence. Empirical evidence supporting this foundational principle of EBM is lacking.

The main purpose of this report is to assess if a) low certainty evidence is more often revised than high certainty evidence in subsequent studies, and if b) the magnitude of effect size differs between high and low CoE.

Methods

We assessed the change in CoE between the original and updated Cochrane systematic reviews, which reported rating of CoE as per the Grading of Recommendations Assessment, Development and Evaluation (GRADE) system for critical appraisal of medical evidence.\(^6\) We used GRADE as this has been widely recognized as the most advanced system for operationalization of fundamental principles of EBM and critical
evaluation of medical evidence. GRADE was developed in the first decade of the 21st Century after critical appraisal of 106 systems for rating the quality of medical research evidence showed that none of them was capable of distinguishing low from high quality evidence. We focused on the assessment of systematic reviews, rather than individual trials, because the second important EBM principle is that assessment of the true effects of health interventions is best accomplished by evaluating total evidence on the topic rather than based on a study selected to favor a particular claim. GRADE is also considered a suitable method to assess certainty of evidence at the level of systematic review/meta-analysis.

Cochrane Reviews are regularly updated providing a unique opportunity to assess when and whether the assessment of CoE changes between the original and updated reviews as a result of new evidence generated between two reviews. Since 2013 Cochrane Reviews have mandated the use of GRADE Summary of Findings (SoF) to summarize CoE and magnitude effects of interventions that the reviews assessed. We evaluated all Cochrane reviews published in the last 5 years in the Cochrane Database of Systematic Reviews [https://www.cochranelibrary.com/cdsr/about-cdsr].

We used SoFs from the original and updated reviews to extract data for the primary outcome related to CoE and to assess the magnitude and direction of effect. (In case of multiple primary outcomes, the data were extracted from the first one listed in SoF table that contained data in both original and updated review). Eligible SR/MAs were divided into 5 groups; data were extracted from each group by pairs of independent reviewers. Kappa interrater agreement was calculated for each pair regarding CoE. As explained, we recorded CoE according to GRADE criteria (very low, low, moderate, high).

We also extracted summary meta-analytic estimates for the primary outcome from each pair of reviews, i.e. point estimates, dispersion (e.g. 95% confidence interval), metric used (e.g., relative risk, odds ratio, hazard ratio, standardized mean differences, etc.), number of trials per meta-analysis, number of participants, type of comparator (active vs placebo/no treatment), type of treatment (pharmaceutical vs non-pharmaceutical), whether the authorship of the original and updated reviews changed (to capture potential differences in judgment of CoE by the review team), and type of studies (randomized controlled trials vs observational studies) that were meta-analyzed.

We converted all effect estimates into odds ratio (OR). We also converted all effect sizes in the same direction, with OR<1 indicating reduction of undesirable outcomes (i.e., more beneficial treatment). Because GRADE separates recommendations as strong vs weak based on the CoE, typically endorsing strong vs weak (conditional) recommendations based on moderate/high vs. low/very low, respectively, our key analysis focused on the differences in effect sizes between these subgroups. We conducted McNemar’s test for paired (before vs after) data to reject the null hypothesis of equal probability that CoE remained the same i.e., in very low/low CoE vs moderate/high CoE groups. To test for linear trend in change of CoE over all categories -from very low to high- we employed a symmetry test with marginal homogeneity tests (which reduces to McNemar’s test for two non-independent categories of observations).

To assess for differences in the magnitude of effect size between original and updated evidence as a function of change in the assessment of CoE we calculated the ratio of odds ratio (ROR) across meta-analytic estimates. ROR compares intervention effects in meta-analysis of trials with very low/low vs those with moderate/high CoE (or vice versa). Thus, if the comparison referred to OR with very low/low vs those with moderate/high CoE pertains to ROR<1, this would mean that treatment effects were more beneficial in meta-analysis of trials with very low/low CoE, while ROR>1 would indicate the opposite. A test of interactions was performed to assess the hypothesis of no difference between the subgroups (i.e. treatments effects in very low/low vs moderate/high CoE). Because of assumed correlations in comparison of treatment effects, we calculated standard errors for ROR by correlating the effect sizes observed in the original vs updated reviews. We obtained the values for correlation coefficients from the data. We performed sensitivity analyses by: a) assuming one correlation coefficient between effects sizes in the original vs updated reviews, and b) calculating correlation coefficients for each subgroup according to direction of treatment effects.
we calculated separate correlation coefficients for the subgroup showing positive, negative and no change in direction of effects between the original vs updated review- three correlation coefficients in total). We also repeated all analyses assuming no correlations between the effect sizes. Since we observed no differences in the results regardless of the postulated assumption, we report the default analysis based on calculation with three different correlation coefficients.

Our hypothesis was that ROR between the subgroups would differ; in addition, we would expect that the effect size would be larger if CoE change from moderate/high to very low/low than other way around.

The analyses were based on using random effect Sidik-Jonkman model. We assessed heterogeneity i.e. dispersion of effect size across the meta-analytic estimates by calculating \( \tau \) (tau) statistic. We used \( I^2 \) statistic to assess inconsistency; \( I^2 \) represents the estimated proportion of the observed variance in true effect sizes across individual meta-analyses rather than sampling error; it depends both on heterogeneity and total variation in the estimates between the analyses.

We complemented assessment of heterogeneity with calculation of the absolute deviation of treatment effects (aROR) as a function of change in CoE. By definition, aROR is positive and reflects the x-fold deviation of treatment effect from OR=1 on the OR scale. Thus, if ROR=0.8 or ROR=1.25, the absolute deviation is equal to aROR=1.25. aROR across all SR/MAs was expressed as (unweighted) median and interquartile range (IQR). We also evaluated how the precision of the estimates changed by calculating the ratio of standard errors for each subgroup summarized as (unweighted) median and IQR. Values > 1 indicate larger standard errors (less precision) associated with given category (e.g., very low/low vs moderate/high) of CoE.

A number of subgroup analyses - all defined a priori and published in the protocol to provide further methodological details - were performed. These include assessment of differences between patient-oriented (e.g., mortality, quality of life etc) vs disease-oriented outcomes (e.g. disease response, laboratory outcomes etc.), effect of a change in authorship between the original and updated reviews, effect of comparator intervention (active treatment vs placebo/no treatment control) and type of treatment category (pharmaceutical vs non-pharmaceutical). Finally, in some cases, the SRs included observational studies along with randomized controlled trials (RCTs) and implausibly large ORs generated in conversion processes from standardized mean differences. We further analyzed these results by performing sensitivity analyses excluding SRs with observational studies and large ORs from the analysis.

This paper is reported per PRISMA guidelines. All analyses were conducted with the Stata,ver17 statistical package.

**Results**

The original search, performed on October 20, 2020, identified 3,323 potentially eligible reviews of which 419 SR were included in the final analysis (Fig 1). Of these, 414(207x2) and 384 (192x2) pairs of the reviews were eligible for the analysis of CoE and effect size, respectively. Total number of trials included in 414 reviews was 4217 (1814 before and 2403 after); mean number of trials per meta-analysis was 10 (minimum: 1, maximum:133). Total number of participants was 3,057,956; mean number of participants per meta-analysis was 10,506 (minimum:16; maximum:1,202,382). Interrater kappa agreement varied from 0.79 to 0.97.

Fig 2 shows comparison of CoE in the original and updated Cochrane reviews across of all categories of CoE (Fig2a) and very low/low to moderate/high (Fig 2b) according to GRADE criteria. Consistent with EBM principles, evidence judged to be of very low/low CoE had 2.1 (1.19 to 4.12; p=0.0065) times higher odds to be upgraded in the future studies than moderate/high CoE (Fig 2b). Similarly, across of all categories of CoE, the test for trend was highly significant, indicating an increased probability of change in CoE from very low to high CoE (p=0.0021 for linear trend). We observed no instance in which high or moderate quality evidence was re-assessed as very low quality evidence in the updated SR, while very low CoE was upgraded to moderate or high CoE in 9/39 of updated SR (Fig 2a).

However, we detected no effect of change in CoE on the magnitude of treatment effects [ROR=1.02 (95%CI:
0.74 to 1.39) for change of CoE from very low/low to moderate/high vs. 1.02 (95%CI: 0.44 to 2.37) for moderate/high to very low/low CoE. Test between the subgroups was not significant (p=1). (Fig 3) Although, as explained earlier, from guidelines recommendations perspectives, GRADE typically groups CoE as moderate/high vs. low/very low, we also tried to compare the effect sizes at the two extremes of CoE: very low vs high. Because we observed no study with high CoE that changed into very low CoE (Fig 2a), ROR was impossible to calculate for this comparison.

Nevertheless, there was larger dispersion in ROR in meta-analyses where CoE changed from moderate/high to very low/low than in the opposite direction. This was probably driven by low power for the analysis instead of the hypothesis that effect size would be larger if CoE changed from moderate/high to very low/low than other way around. We had half as many of meta-analyses available for the assessment of ROR based on change of CoE from moderate/high to very low/low (n=16) as those in which CoE changed from very low/low to moderate high (n=33).

aROR was similar between the subgroups [median (IQR):1.12 (1.07 to 1.57) vs 1.21 (1.12 to 2.43)] (Fig 4a, Table 1). As in case of ROR, we observed larger dispersion in aROR in meta-analyses where CoE changed from moderate/high to very low/low than in the opposite direction (Fig 4 a, Fig 4b).

The meta-analyses with no change in CoE had similar ROR [ROR=1.01 (95%CI: 0.85 to 1.21)] (Fig 3b) and aROR [median (IQR):1.13 (1.04 to 1.66)] (Table 1, App Fig 4 and App Fig 4a) to those MAs in which CoE changed (Fig 4 and App Fig 4a). Inconsistency was large across of all meta-analytic estimates ($I^2=99\%$). Likewise, the ratio of standard errors was > 1 [median: 1.09; IQR:0.72 to 1.46] indicating imprecision in the estimates.

Qualitative analysis indicated that direction of the effect changed in 6 SR/MAs only: two in the reviews in which CoE changed from very low/low to moderate/high (of which one was statistically significant ) and in 4 SR/Mas with no change in the assessment of CoE (of which one was statistically significant) (Fig 5, App Figs 12 and 13).

Sensitivity analyses for all defined subgroups showed no change in the results. In fact, when non-randomized studies or outliers were excluded from the analyses, no statistically significant changes were seen in any of the analyses (Appendix).

Discussion

Almost 30 years ago, EBM\cite{23} was introduced to wide medical audience, subsequently being assessed to represent one of the most important medical milestones of the last 160 years, in the same category as innovations such as antibiotics and anesthesia.\cite{24} At the heart of EBM is notion that “not all evidence is created equal”- some evidence is more credible than others; the higher quality of evidence, the more accurate and trustworthy are our estimates about true effects of health interventions.\cite{1} Surprisingly, however, the relationship between CoE and estimates of treatment effects has not been empirically evaluated.

Here, we provide the first empirical support for the foundational EBM principle that low-quality evidence changes more often than high CoE (Fig 2). However, we found no difference in effect sizes between studies appraised as very low vs high CoE (or, very low/low vs. moderate/high CoE (Fig 3)]. This implies that effects that are assessed as less trustworthy/potentially unreliable (as when CoE is low) cannot be distinguished from those assessments, which are presumably more trustworthy/accurate (as when CoE is high). If the magnitude of treatment effects cannot be meaningfully distinguished from evidence appraised as high vs. low quality, then the core principle of EBM seems to be challenged.

Our “negative” results should not be construed as a challenge to sound, normative EBM epistemological principles, which hold that optimal practice of medicine requires explicit and conscientious attention to the nature of medical evidence.\cite{1,25,26} Rather, in assessing the relationship between CoE and “true” effects of health interventions, more salient question is to ask if the current appraisal methods capture CoE as intended by the EBM principles. Critical appraisal of CoE is integral aspect of conduct of systematic reviews, guidelines development and is widely accepted in the curricula in most medical and allied professional schools
across the world. Over the years, many critical appraisal methods have been developed\(^1\) to eventually culminate in development of GRADE methodology, which has been endorsed by more than 110 professional organizations.\(^7\) However, as we demonstrate here, despite GRADE’s capacity to distinguish across CoE categories, it could not- and we suspect none of other appraisal methods that GRADE has replaced- reliably discerned the influence of CoE on the estimates of treatment effects. The results agree with those of Gartlehner et al who, based on cumulative meta-analysis of 37 Cochrane reviews, found\(^27\) limited value of GRADE in predicting stability of strength of evidence as new studies emerged.

The finding that the magnitude of effect size is not reflected in a change of CoE is surprising as previous meta-epidemiological studies showed that various study limitations that affect CoE significantly influence estimates of treatment effects\(^28\) (although not always consistently\(^16\)). For example, as measured by ROR, inadequate or unclear (vs. adequate) random-sequence generation, inadequate or unclear (vs. adequate) allocation concealment, or lack of or unclear double-blinding (vs. double-blinding) led to statistically significant exaggeration of treatment effects by 11%, 7% and 13%, respectively.\(^28\) These study limitations are taken into account in rating of CoE using GRADE method\(^6\), so one would expect that effect size would differ between low vs high CoE in the GRADE assessment. However, on further examination, we observe that GRADE combines the study limitations such as adequacy of allocation concealment, blinding etc (risk of bias) with the assessment of inconsistency, imprecision, indirectness and publication bias to assign the final rating of CoE (from very low to high quality) in additive fashion.\(^12,29\) It appears that using additive means to report the properties of negative and positive changes in treatment effect could unhelpfully neutralize this effect and cause imprecision in the overall estimate. Thus, one can have the same estimates of treatment effects but completely different GRADE ratings. This is, however, problematic because central assumption of GRADE is that estimates underpinned by high CoE are unlikely to change, whereas the very low/low CoE estimates are more likely to change.

A potential limitation of our study is that we have not collected data on the individual factors that drove assessment of CoE (i.e., study limitations/risk of bias vs inconsistency, imprecision, or indirectness, for example). However, the present empirical report targets, for first time, the end-stage level assessment of CoE, according to GRADE specifications, which is how CoE is used in practice to aid interpretation of evidence and affect development of clinical guidelines.

We also detected imprecision in the estimates of effects sizes and relatively wide ROR confidence intervals, particularly in the subgroup of meta-analyses describing treatment effects in the reviews with CoE that changed from moderate/high to low/very low. It may be argued that the current methods of CoE appraisal are simply not sensitive enough and that with much larger sample size of SR/MA, we would be able to differentiate between effect sizes across categories of CoE. This point was made by Howick and colleagues\(^30\) who showed no change in the CoE between original and updated reviews in a set of the 48 trials they examined, albeit they made no attempt to identify changes in effect sizes. However, obtaining larger sample sizes is unrealistic given that we reviewed almost all SRs in the Cochrane database since the GRADE assessment of CoE was mandated (up to May 2021). Finally, few Cochrane Reviews we analyzed included observational studies. It is possible that GRADE may not differentiate the quality of randomized evidence well but that it may perform better if the comparison is made between randomized vs observational studies. The Cochrane Reviews, however, are typically based on randomized trials. Therefore, categorization of CoE based on currently mandated critical appraisal system using GRADE in the Cochrane Reviews does not meaningfully separate effect sizes across the existing gradation of CoE (although, capacity of GRADE to distinguish the magnitude of effect size between randomized and observational studies outside of the purview of Cochrane Reviews remains a worthwhile goal for further empirical research).

Given that studies can be well done, and correctly estimated treatment effects, but be poorly reported\(^31,32\), it is also possible that we could not detect influence of CoE on the estimates of treatment effects because current critical appraisal methods depend on the quality of reporting of the trials that are selected for meta-analysis. However, if we believe that quality of reporting does not matter, then the entire critical appraisal efforts can be considered misplaced to begin with.
Conclusions

To the extent that the central to the epistemology of EBM is that what is justifiable or reasonable to believe depends on CoE\(^1\), our findings indicate urgent need to refine current EBM critical appraisal methods. If EBM is going to flourish, it is crucial to develop methods with capacity to categorize CoE to reliably differentiate between magnitude effects that are potentially biased from those that are accurate and trustworthy. The major opportunity, therein, lies in addressing the main limitations of this study—carefully and painstakingly discerning various aspects of CoE (from the components related to study limitations/risk of bias to inconsistency, imprecision, or indirectness) to better characterize CoE and its relationship to the magnitude of effects of health interventions.

Contributors and sources

The authors are notable as an interdisciplinary team of EBM practitioners and instructors who are respected as clinicians, mathematicians, epidemiologists, statisticians, methodologists, and researchers across academic institutions, hospitals and clinics in the UK, Can, USA, Brazil and Switzerland. Their research experience ranges from recently acquired doctorates to over 40 years in research and clinical practice. All authors contributed to the methods, commented on the analysis and contributed to writing and revising the manuscript. Our sources and selection criteria are contained within the document, the data is publicly available from the Cochrane Database and our statistical methods are outlined in the methods, figures and tables. PRISMA was used to report our findings. BD serves as the guarantor of the article.

A conceptual idea: BD

Design: BD & DN

Protocol development: BD, MMA, DN, LH, DK, AP, RR, PN, APPdS, DM, RLP, LEF

Data acquisition: MMA, DK, AP, RR, PN, APPdS, DM, RLP, LEF, RP

Statistical analysis: IH, BD, LH

Drafting manuscript: BD

Critical revision of the manuscript for important intellectual content: BD, LH, DN, AP, DK, RR, PN, APPdS, DM, RLP, LEF, RP

Administrative, technical, or material support: BD, MMA

Supervision: BD

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Conflict of interest: We declare no conflict of interest in relation to this work

Legends:

Fig 1 PRISMA diagram (study flow diagram for evidence source and selection)

Fig 2 Change in certainty of evidence (CoE) in original and updated Cochrane systematic review. A) across all categories of CoE as characterized by GRADE; b) grouped as very low/low vs moderate/high quality evidence

Fig 3 Comparison of effects of health interventions in meta-analyses in which certainty of evidence (CoE) changed from very low/low to moderate/high vs effects in meta-analyses where CoE changed from modere-
ate/high to very low/low (a); b) summary of studies shown in a) with addition of comparison of meta-analyses where CoE did not change. Abbreviations: ROR-ratio of odds ratio; $\tau^2$ statistic and $H^2$ – measures of heterogeneity; $I^2$ statistic-measure of inconsistency.

Fig 4. a) Absolute deviation (AD) of treatment effects (aROR) in meta-analyses in which certainty of evidence (CoE) changed from very low/low to moderate/high vs effects in meta-analyses where CoE changed from moderate/high to very low/low; b) summary of aROR by change in CoE [For graph displaying aROR for all studies, including those that did not have change in CoE, see Appendix, App Fig 4 and App 4a]

Fig 5 Change in effect size, qualitative analysis (see also App Figs 12 and 13).

References


22. STATA, ver. 17 [computer program]. College Station, TX2021.


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Fig 2

a)  

b)  

Fig 3

a)  

b)  

d) ORs by change in QoE

Study  OR  95% CI  p-value

Grown_GoE change

Very/Low, Low = Med/High  1.92 [1.46, 2.47]  0.001

Med/High, Very/Low, Low = 1.82 [1.24, 2.67]  0.001

Ditch change

1.91 [1.85, 2.00]  0.000

Test of group differences: G2=1.90, p=0.18

Changed_GoE

No change  141

1.81 [1.33, 2.48]  0.001

Change_GoE  47

1.80 [1.34, 2.42]  0.001

Test of group differences: G2=1.12, p=0.30

Overall

Heterogeneity: I^2 = 71.5%, Q=63.41, p<0.000

Test of K=10, Q(10)=14.40, p=0.03

Random-effects Sib-Jenkins model
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Table 1 Summary of aROR (absolute deviation) by change in QoE

<table>
<thead>
<tr>
<th>Study</th>
<th>K</th>
<th>aROR</th>
<th>95% CI</th>
<th>p-value</th>
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</thead>
<tbody>
<tr>
<td>&quot;Moderate&quot; change</td>
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<td>1.40 [1.07, 1.80]</td>
<td>0.046</td>
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<tr>
<td>&quot;Higher&quot; -- &quot;Lower&quot;</td>
<td>15</td>
<td>1.80 [0.70, 3.57]</td>
<td>0.008</td>
<td></td>
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<tr>
<td>&quot;Lower&quot; change</td>
<td></td>
<td>1.47 [1.11, 1.98]</td>
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<tr>
<td>&quot;Higher&quot; change</td>
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<td>1.47 [1.11, 1.98]</td>
<td>0.007</td>
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<tr>
<td>Overall</td>
<td></td>
<td>1.47 [1.11, 1.98]</td>
<td>0.007</td>
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</tr>
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</table>

Random effects (Snedecor model)